Candidate genes versus genome-wide associations: which are better for detecting genetic susceptibility to infectious disease?

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Technological developments allow increasing numbers of markers to be deployed in case-control studies searching for genetic factors that influence disease susceptibility. However, with vast numbers of markers, true ‘hits’ may become lost in a sea of false positives. This problem may be particularly acute for infectious diseases, where the control group may contain unexposed individuals with susceptible genotypes. To explore this effect, we used a series of stochastic simulations to model a scenario based loosely on bovine tuberculosis. We find that a candidate gene approach tends to have greater statistical power than studies that use large numbers of single nucleotide polymorphisms (SNPs) in genome-wide association tests, almost regardless of the number of SNPs deployed. Both approaches struggle to detect genetic effects when these are either weak or if an appreciable proportion of individuals are unexposed to the disease when modest sample sizes (250 each of cases and controls) are used, but these issues are largely mitigated if sample sizes can be increased to 2000 or more of each class. We conclude that the power of any genotype–phenotype association test will be improved if the sampling strategy takes account of exposure heterogeneity, though this is not necessarily easy to do.

Keywords: tuberculosis; association study; power; infectious disease; heterozygosity; population simulation

1. INTRODUCTION

Understanding the genetic basis of susceptibility to disease has become an increasingly important target for research. Moreover, with the publication of complete genome sequences for humans and other species, the development of literally millions of polymorphic markers [1, 2] and emerging technologies that make it feasible to collect genetic data at unprecedented rates, we appear to be faced with an embarrassment of riches. However, progress remains slower than many might predict [3], raising the question of why this power is not more effective.

Studies aimed at uncovering genetic predispositions to disease usually attempt to demonstrate an association between the genotype at one or more polymorphic markers and a phenotype related to disease susceptibility [4]. There are two main approaches, one based on candidate genes (CG) [5, 6], the other based on testing the entire genome (genome-wide association (GWA)). Both approaches enjoy a combination of benefits and drawbacks. Broadly, CG studies tend to have rather high statistical power but are incapable of discovering new genes or gene combinations, while GWA studies can pinpoint genes regardless of whether their function was known before [7] but have low power owing to the number of independent tests performed [3, 8, 9]. Indeed, the problem of false positives, already an issue in early studies deploying a few hundred microsatellite markers, is becoming acute as we move into the era when single nucleotide polymorphisms (SNPs) are replacing microsatellites and more than a million markers may be used.

With one to two million SNPs it is estimated that an alpha level of $5 \times 10^{-8}$ is required [3], though this number may be reduced by using a subset of ‘tag SNPs’ whose genotypes correlate strongly with those of neighbouring loci [10, 11].

Recent reviews have reached contrasting conclusions about the effectiveness of CG compared with GWA approaches. One analysis asks whether GWA studies ‘are a waste of time’ since even with thousands of samples the approach can be underpowered [12]. However, SNP-based GWA studies in humans have had considerable success in identifying regions that are important in disease development in such conditions as diabetes [13], Crohn’s disease [14] and other autoimmune and genetic conditions. By contrast, GWA studies have had relatively limited success in finding novel genes involved in susceptibility to infectious diseases. One reason for this may lie with a number of added complexities that infectious diseases imply, including the key element of exposure. Only exposed individuals can contract the disease, meaning that even highly susceptible genotypes can be found in unaffected individuals. Moreover, exposure is often correlated with other factors such as age and behaviour.

Issues such as exposure imply studies of genetic susceptibility to infectious disease will have modest effect sizes and hence that statistical power will be at a premium. If so, then the greater inherent power of CG may make it preferable over GWA. This expectation seems partially fulfilled, in that small-scale studies in both human [15] and non-human systems have proved surprisingly effective in identifying genomic regions associated with disease susceptibility [16–18]. The non-human studies are particularly interesting because as few as around 10 markers have revealed rather convincing
links to immune-related genes such as methylmalonyl-CoA [19,20], despite often not being initially selected for proximity to CGs. This apparent success probably reflects more than just the reduced need for multiple testing. Thus, microsatellites have higher variability, increasing the chance that one allele shows strong linkage disequilibrium with a causative gene allele; microsatellites are selected for high polymorphism, increasing the chance that such markers lie near genes experiencing balancing selection [21]; genetic effects in natural populations may be much stronger than in humans owing to the absence of medical intervention.

The problems associated with finding genes that influence susceptibility to infectious diseases are nicely illustrated by tuberculosis, an important disease caused by various species of Mycobacteria and affecting many species including humans and cattle [22,23]. In the UK cattle herd, bovine tuberculosis (bTB) is monitored using a skin test that is less than 100 per cent effective [24]. Cows testing positive (‘reactors’) may variously have a current infection, a latent infection, have previously resisted infection or were infected but are now recovered. Exposure rates also vary widely with factors such as farm management practice [25], proximity to infected badgers [26] and geography. Untangling the influence of genetics versus exposure is difficult. Thus, a farm may lack disease because it has no exposure or because its (related) cattle mostly carry an allele that confers sufficient resistance. Conversely, farms may have high incidence because of high exposure or high genetic susceptibility.

Attempts to calculate statistical power in association studies in order to determine the optimum sample size for individuals and markers have generated a large literature [11,27,28]. However, analytical calculations of power necessarily have to ignore a range of real, stochastic complexities such as variation in exposure rate, variation in recombination rate, the impact of local selection pressure, population structure and the genetic background of the original causative mutation. Consequently, for human studies, the tendency is to assume the use of SNPs and then to exploit the wealth of HapMap data to estimate parameters such as average levels of linkage disequilibrium [10,29,30]. However, for non-human systems and microsatellite data, these calculations have little relevance. Moreover, most power calculations assume that the most resistant genotype is homozygous [28,31], as do methods used to correct for population structure [32], when for infectious diseases it may well be the heterozygote that is most resistant. Thus, current methods for calculating power appear poorly suited to studies of infectious disease, particularly in non-human systems.

In view of the above, we decided to conduct a simulation-based study to assess the relative power of GWA and CG approaches to detect genetic susceptibility to infectious disease in the presence of varying exposure rates. We consider both microsatellite and SNP markers and explore the relative chance that an informative marker lies near to a susceptibility factor or is, in the case of a SNP, the functional mutation itself. Although inspired by studies focused on bTB, our approach has implications for any case–control studies where factors such as variable disease exposure cause individuals with susceptible genotypes to be included in the control group.

2. RATIONALE AND METHODS

We consider a single gene affecting the probability that an individual exposed to a pathogen contracts the disease. Individuals are deemed either ‘susceptible’ or ‘resistant’ and may be either ‘exposed’ or ‘unexposed’ to the disease. Unexposed individuals do not get the disease while exposed susceptible individuals always do. Exposed resistant individuals become infected with a probability that can be varied to simulate different levels of genetic benefit. We refer to this as the resistant infected fraction (RIF). Naturally, susceptible individuals may not be 100 per cent susceptible, but in our model, such individuals are treated as unexposed: effectively, exposure is normalized to the point where all susceptible individuals become infected.

Genetic susceptibility is assumed to be dictated by a single gene with two alleles where heterozygotes are resistant and homozygotes are susceptible. Other scenarios are possible, but are difficult to simulate because any directional selection (i.e. the disease influences survival/ reproduction) tends to lead to the elimination of susceptible alleles. Near the gene is a single locus, designated the marker, that mutates according to a strict stepwise mutation model. By varying the mutation rate, it is possible to select scenarios where the population carries three genotypes, taken as equivalent in informativeness to a SNP, or more than 10 genotypes, taken as representing a microsatellite. In each simulation, a population of 1000 diploid individuals is initialized with a single marker allele and random gene alleles (0 and 1, each with \( p = 0.5 \)). Evolution progresses under random mating, selection (heterozygote fitness = 1, homozygote fitness = 0.75), mutation at the marker and recombination (range examined \( 10^{-3} - 10^{-5} \) per generation).

At the end of each simulation, the frequencies of each compound genotype (gene and marker) are determined and, within each, the expected frequency of diseased animals determined according to the risk factors of gene genotype, exposure and RIF. Based on these frequencies, a sample of \( N \) cases and \( N \) controls is generated deterministically and a simple \( \chi^2 \)-test used to test for a significant difference in genotype frequencies between cases and controls. Using the same terminal genotypes, this process is then repeated across a range of exposures (range tested \( = 0.1 - 0.9 \) in steps of 0.1) and a range of RIF values (range tested \( = 0.1 - 0.9 \) in steps of 0.1). Simulations are repeated 100 times for each recombination rate, yielding a per cent score for how often the \( \chi^2 \)-test yielded a significant association. For significance, we use an alpha level equivalent to \( p = 0.05 \) after correction for the number of markers tested, i.e. with \( X \) markers, we use \( \alpha = 1/20X \). Finally, to simulate a causative SNP in the gene itself, the same process is repeated, but this time using the genotype of the gene instead of the marker. For GWA we assumed 50,000 SNPs, while for the CG study we assume 10 candidate markers. The stochastic simulation program was coded in C++, using Marsenne Twister random number generation.

Using many SNPs invokes an important trade-off. Using more SNPs increases the chance that the functional SNP is included, but reduces power owing to an increase in false positives. To obtain a measure of whether there is an optimal number of SNPs, we simulated this scenario. We assumed a maximum of five million SNPs in the

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Proc. R. Soc. B (2011)
Candidate genes versus genome-wide tests  W. Amos et al.  1185

Figure 1. How the probability of finding a significant genotype–phenotype association for an infectious disease varies with marker type, exposure rate, recombination rate and size of genetic effect. Each cell contains the proportion of approximately 100 replicate simulations that yielded an experiment-wide significant association at \( \alpha = 0.05 \). Greyscale represents: white = 0–5% significant; light grey = 5–50% significant; dark grey = 50–90% significant; black = 90–100% significant. Exposure is the proportion of individuals exposed to the disease, genetic effect is the probability of an exposed, genetically resistant individual catching the disease relative to a susceptible individual (\( p = 1 \)). (a) ‘SNPs’ (= marker carries three genotypes), recombination rate = \( 10^{-3} \); (b) ‘SNPs’, recombination rate = \( 10^{-5} \); (c) ‘microsatellites’ (= marker carries 10+ genotypes), recombination rate = \( 10^{-3} \); (d) ‘microsatellites’, recombination rate = \( 10^{-5} \).

The genome and that a study uses somewhere between 2000 and 2000 000 randomly selected SNPs. The probability of detecting an effect is then taken as \( 1 - (1 - P)(1 - Q) \), where \( P \) is the probability that the functional SNP is included in the marker set, \( Q \) is the probability that the functional SNP itself detects a significant association at alpha (see above), \( R \) is the probability that at least one SNP is close enough to the gene to be linked by a recombination rate of \( 10^{-3} \), and \( S \) is the probability that a linked SNP detects a significant association. Since cattle exhibit high levels of linkage disequilibrium owing to their small effective population size \[33\], we conservatively equate a recombination rate of \( 10^{-5} \) with a distance of 0.1 Mb.

3. RESULTS

Figure 1 summarizes the results of our simulations for the two extreme recombination rates, \( 10^{-3} \) and \( 10^{-5} \), and the two marker categories, SNPs (= three genotypes) and microsatellites (10+ genotypes). Several trends are apparent. First, as expected, higher exposure rates and higher effect sizes (difference in susceptibility between susceptible and resistant genotypes) both act to increase the chance that a significant association is detected. Noticeably, at exposure rates below around 30 per cent only very strong genetic effects (twofold or greater risk) are likely to be detected. Second, low recombination rates promote detection, with the \( 10^{-3} \) level requiring almost deterministic infection where genetic resistance is higher than 70 per cent. Third, where associations are detected, microsatellites tend to be around twice as likely as a SNP to detect an association, the exception being around the borderline of detection, where SNPs perform somewhat better. We interpret this as reflecting the \( \chi^2 \)-test; for medium-strength associations, the high genotype diversity helps increases power, but when the effect is weak, the loss of degrees of freedom becomes critical.

Figure 2 summarizes the results of a simulated GW A study in which we assume that 50 000 SNPs are deployed and include the functional mutation itself. Here, a rather similar region of the parameter space is filled with zeros compared with the \( 10^{-5} \) recombination rate simulations. By implication, when the genetic effect is weak and/or the exposure rate is low, throwing more SNPs at the problem has relatively little impact on the chance of finding a significant association. However, almost all the non-zero cells contain ‘1’s, indicating that where the approach can detect an association, it almost invariably does so. Compared with linked markers, there are few grey areas where some studies will find an effect and others not.

Finally, we considered the impact of using larger sample sizes and the way the joint probability that a SNP is either in strong linkage disequilibrium with the functional gene or is the gene mutation itself varies with SNP number. Figure 2b summarizes results for a CG approach with 10 microsatellites and a sample size increased from 500 to 4000. Only the weakest genetic effects are not detected with high probability and exposure rates have reduced influence. Figure 3 summarizes the results for simulations based either on 50 000 or two million SNPs for the two different sample sizes. Increasing the number of SNPs to two million brings some benefit, though arguably not commensurate with the added empirical effort. As with the CG approach, increasing sample sizes brings much greater benefit, with a much larger proportion of scenarios yielding good to high power. Noticeably, for an equivalent sample size, the CG approach generally offers higher power.

4. DISCUSSION

We have used stochastic simulations to explore how the power of genetic association studies to detect genes that influence susceptibility to an infectious disease vary with factors such as exposure rates and the strength of genetic
resistance/susceptibility. We find that when either exposure rates are very high or the genetic effects are strong, an association can be detected, but that many plausible scenarios are unlikely to be detected regardless of the experimental effort deployed. In our simulations, a CG approach tends to outperform a GWA approach, even when this involves literally millions of SNPs.

GWA studies are expensive and are currently outside the likely budget for most non-model species study systems. However, rapidly improving technology and falling prices mean that we are just reaching the point indeed starting [34]. Consequently, it seems a good time to ask whether such investment is likely to yield results commensurate with the outlay, a question whose importance is emphasized by recent papers that suggest the likely budget for most non-model species study systems. The above contrast between high and low breakdown farms illustrates two important points. First, resistance and susceptibility are not simply the opposite sides of the same coin. From a genetic perspective, what matters is which genotypes are common and which are rare: for example, if most of the genetic effect is manifest as resistance/susceptibility. We find that when either exposure rates are very high or the genetic effects are strong, an association can be detected, but that many plausible scenarios are unlikely to be detected regardless of the experimental effort deployed. In our simulations, a CG approach tends to outperform a GWA approach, even when this involves literally millions of SNPs.

Comparing the CG and GWA approaches suggests a general advantage to CG, regardless of the number of SNPs being used. Several effects contribute to this overall picture. With a few thousand SNPs, the problem of false positives is small but so is the chance that one of the SNPs used is close to a particular gene. Moreover, even if a SNP is close enough, the higher diversity of microsatellites appears, under most circumstances, to increase the
chance of finding an association, presumably because more alleles make it more likely that one is strongly indicative of a key allele at the gene. With medium numbers of SNPs, say 50,000, strong linkage disequilibrium with any given gene becomes likely, but correction for multiple testing reduces the power of the test, handing the advantage to a method that chooses the gene \textit{a priori}. Finally, with a million SNPs or more, it becomes increasingly likely that one SNP is the actual functional mutation. When this happens, the effect size leaps and extremely small \(p\)-values tend to result. However, the probability of finding an experiment-wide significant association remains well below one because, unless all possible SNPs are deployed, every time the functional SNP is not included, the signal from the gene itself gets lost in a veritable sea of false positives.

To say that the CG approach is superior would overstate the case. If the list of target genes identified \textit{a priori} fails to include the gene(s) that are important, this method will also fail. For example, an individual might become susceptible indirectly, through being in a poor nutritional state owing to a gene involved in metabolism, and such genes are unlikely to be included as candidates. Second, we have assumed a modest sample size of 500 samples. Increasing sample size increases statistical confidence, reducing the impact of false positives and making any given association more detectable. We tested this by repeating the simulations with 2000 cases and 2000 controls, finding an increase in the parameter space in which associations could be detected, with exposure rates having a reduced impact and only the weakest genetic effects (less than 30% fitness advantage) likely to pass undetected. Despite this, the relative performances of CG and GWA remain rather unchanged, with any given combination of parameter values being more likely to be detected by CG. This presumably reflects the fact that, if the primary functional gene is included as a candidate for CG, GWA studies only ever have an advantage when they include the functional mutation itself, and this depends on SNP number not sample size.

In conclusion, variable exposure rates present an important challenge to the experimental design of association studies. Greatest statistical power is gained by sampling exposed individuals across a range of genotypes, including both those that are relatively resistant and those that are susceptible. However, it is potentially dangerous to base a sampling strategy primarily on disease incidence assuming this to be a proxy for exposure, because the more this assumption is true, the weaker, and hence less detectable, the genetic effects will be. Our simulations suggest that CG approaches tend to offer some extra power compared with genome-wide tests, though this assumes prior knowledge of genes likely to play a role. This power gain is most likely to be important in smaller studies, since when thousands of samples can be analysed, all but weak genetic effects are likely to be detected by both approaches. Our work thus supports the handful of recent papers that question whether very large SNP-based studies offer the only viable approach [12], and instead suggests that much cheaper analyses based on microsatellites can be highly effective.

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