Thermal conductance and basal metabolic rate are part of a coordinated system for heat transfer regulation

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Thermal conductance measures the ease with which heat leaves or enters an organism’s body. Although the analysis of this physiological variable in relation to climatic and ecological factors can be traced to studies by Scholander and colleagues, only small advances have occurred ever since. Here, we analyse the relationship between minimal thermal conductance estimated during summer (Cmin) and several ecological, climatic and geographical factors for 127 rodent species, in order to identify the exogenous factors that have potentially affected the evolution of thermal conductance. In addition, we evaluate whether there is compensation between Cmin and basal metabolic rate (BMR)—in such a way that a scale-invariant ratio between both variables is equal to one—as could be expected from the Scholander–Irving model of heat transfer. Our major findings are (i) annual mean temperature is the best single predictor of mass-independent Cmin. (ii) After controlling for the effect of body mass, there is a strong positive correlation between log10 (Cmin) and log10 (BMR). Further, the slope of this correlation is close to one, indicating an almost perfect compensation between both physiological variables. (iii) Structural equation modelling indicated that Cmin values are adjusted to BMR values and not the other way around. Thus, our results strongly suggest that BMR and thermal conductance integrate a coordinated system for heat regulation in endothermic animals and that summer conductance values are adjusted (in an evolutionary sense) to track changes in BMRs.

1. Introduction

Thermal conductance measures the ease with which heat leaves or enters an organism’s body (see the electronic supplementary material, box S1) [1]. Although the analysis of this variable in relation to climatic and ecological factors can be traced to Scholander et al. [2–4] back to the Fifties, only small advances have been made ever since those pioneering works. For instance, according to McNab’s recent review [5], very few articles dealing with variability and macrophysiological patterns in thermal conductance have been published over the last three decades.

In comparative studies, minimal thermal conductance (Cmin) is considered more informative than thermal conductance because in the former, the external factors that modify heat exchange have been eliminated and also because behaviours affecting thermal conductance have been standardized (electronic supplementary material, box S1) [1]. An interesting point regarding Cmin is its relationship with basal metabolic rate (BMR). According to the Scholander–Irving classical model of heat transfer, the critical temperature differential (ΔTm, °C) between body (Tb, °C) and ambient temperature (Ta, °C) is equal to BMR/Cmin (°C); which means that both lower Cmin and higher BMR may contribute to an increase in
thermoregulatory capacity by allowing heat conservation and heat generation, respectively (see electronic supplementary material, box S1) [1,3]. In line with this, McNab [5] recently pointed out that ‘An appreciable decrease in conductance in some cases appears to compensate for a reduction in the rate of metabolism (p. 19) ...which emphasize that rate of metabolism and thermal conductance are part of a coordinated system (p. 20)’. However, as far as we know, few empirical tests have been conducted to evaluate this intriguing statement. For instance, compensation in \( C_{\text{min}} \) values has been suggested for some rodent species inhabiting xeric habitats [6–8] and also for a few Carnivora species with frugivorous or mixed diets [9].

Within this conceptual framework, the aims of this study were twofold. First, to explore the relationship between \( C_{\text{min}} \) and several ecological, climatic, and geographical variables for 127 rodent species, in order to identify the exogenous factors that have potentially affected the evolution of thermal conductance. Second, to explicitly test whether there is compensation between \( C_{\text{min}} \) and BMR in such a way that the scale-invariant ratio between \( \log_{10}(C_{\text{min}}) \) and \( \log_{10}(\text{BMR}) \) is equal to one.

2. Material and methods

(a) Database description

We downloaded data on minimal thermal conductance estimated during summer months (\( C_{\text{min}} \)), body mass (\( m_{b} \)), BMR and geographical coordinates (together with altitude) for 127 rodent species, compiled by Lovegrove [10] (see the electronic supplementary material, table S1). Methodological details on how these variables were assessed are explained elsewhere [10]. Then, for each data point downloaded, we obtained from WorldClim database (http://www.worldclim.org/) the following climatic variables: annual mean temperature (\( T_{\text{med}} \) in °C), minimum temperature of the coldest month (\( T_{\text{max}} \) in °C), maximum temperature of the warmest month (\( T_{\text{max}} \) in °C), temperature amplitude and temperature seasonal amplitude, mean monthly temperature range (MMTR: mean of the maximum temperature minus mean of the minimum temperature, in °C), temperature seasonality (TS: standard deviation of the mean monthly temperature, in °C), mean monthly temperature range (MMTR: mean of the maximum temperature minus mean of the minimum temperature, in °C), isothermality (IT: the mean monthly temperature range divided by the temperature annual range, thus adimensional), accumulated annual rainfall (rainfall, in mm), rainfall of the driest month (\( R_{\text{min}} \), in mm), rainfall of the wettest month (\( R_{\text{max}} \), in mm) and rainfall seasonality (RS: standard deviation of the mean monthly rainfall, in mm) (see the electronic supplementary material, table S1). These variables were obtained using the free software Diva-Gis (http://www.diva-gis.org/). In addition, a net primary productivity map (based on [11]) was downloaded from Socioeconomic Data and Application Center homepage (http://sedac.ciesin.columbia.edu/es/hanpp.html) and an aridity map (based on [12]) was downloaded from the CGIAR Consortium for Spatial Information homepage (http://csi.cgiar.org/Aridity/). Net primary productivity (NPP, in tons of carbon per 0.25° of latitude cell) and aridity index values (\( A\text{ridity} \), adimensional) were then obtained for each site using the software ArcGis v. 10 (see the electronic supplementary material, table S1). Aridity index values were multiplied by \(-1\) in order to obtain a direct relationship between real aridity and index values; thus, in our scale, values larger than \(-0.030\) represent hyper-arid environments, values between \(-0.031\) and \(-0.200\) represent arid environments, values between \(-0.201\) and \(-0.500\) represent semi-arid environments, values between \(-0.501\) and \(-0.650\) represent dry sub-humid environments and values lower than \(-0.651\) represent humid environments. Finally, we compiled data on rodent food habits from the literature and assigned each species to one of the following dietary categories: herbivorous (H), herbivorous–granivorous (HG), granivorous (G), omnivorous (O) and insectivorous (I) (see the electronic supplementary material, table S1).

(b) Data analysis: exogenous factors affecting \( C_{\text{min}} \)

The relationships between \( C_{\text{min}} \) and exogenous factors were evaluated through standard least-squares regression techniques, using body mass as a covariate. A \( \log_{10} \) transformation was applied to \( C_{\text{min}} \) and \( m_{b} \) raw data in order to meet the assumption of normality. In these regression analyses, species food habits were ranked according to their approximate assimilable energy content [13] and included as an ordinal variable (herbivorous: 1, herbivorous–granivorous: 2, granivorous: 3, omnivorous: 4 and insectivorous: 5). We estimated the goodness of fit of the 57,343 possible models without interaction terms—i.e. combination of 16 independent variables but taking just two of the following three variables: \( T_{\text{min}} \), \( T_{\text{max}} \) and TAR at a time (because TAR is equal to \( T_{\text{max}} \) minus \( T_{\text{min}} \))—and used the Bayesian Information Criterion (BIC) to compare them. Specifically, a model was selected as a ‘good model’ if its BIC value did not differ from the overall best model BIC value (which represents, by definition, the lowest BIC value) in more than 2.3 units [14]. All these analyses were performed using the R package leaps [15].

For models selected as ‘good models’, we evaluated the effect of phylogeny on the relationship between \( C_{\text{min}} \) and exogenous
Table 1. Parameter estimation (and s.e.) for models including only one exogenous factor in addition to body mass \((m_b)\). BIC, Bayesian Information Criterion value; \(\Delta \text{BIC}\), BIC model – lowest BIC; \(r^2\) = proportion of variance explained by the model. The \(\log_{10}\) of \(C_{\text{min}}\) and \(m_b\) were used in these analyses. The probability value associated with each exogenous factor is denoted by asterisks. See material and methods for factors abbreviations.

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\* \(p < 0.1\), ** \(p < 0.05\), *** \(p < 0.01\).

Factors, using a Bayesian Phylogenetic Mixed Model (Bayesian PMM) [16,17], in addition to Bayesian Model Averaging (BMA) [14]. The phylogenetic tree published by Lovegrove [10] was transformed to a Newick formatted tree using the program TreeSnatcherPlus [18]. Taking this tree as the starting point, we decided to incorporate phylogenetic uncertainty in the calculations using BMA because: (i) branch lengths are not known for this tree, and (ii) there are several soft polytomies associated with (i). In consequence, phylogenetic uncertainty was included by generating 1000 trees in which polytomies were randomly resolved (by transforming all multi-chotomies into a series of dichotomies with one or several branches of length zero) and branch lengths were randomly sampled from a uniform distribution (ranging between 0.01 and the maximum branch length). Inverse-Wishart distributions were used as prior for variances (scale = 2, d.f. = 2) with 350,000 iterations, 150,000 of burn-in and a thinning interval of 20. For each comparative model, the effect of exogenous factors on \(C_{\text{min}}\) was calculated through linear mixed models, using body mass as a covariate. Then, to estimate the effect of each exogenous factor on \(C_{\text{min}}\), we calculated the proportion of posterior estimates greater than zero \((g > 0)\). In short, \(g > 0\) can be viewed as the probability of observing either a positive (if \(g > 0.5\)) or negative (if \(g < 0.5\)) association between the dependent variable (i.e. \(C_{\text{min}}\)) and each exogenous factor. Note that when the dependent variable is not affected by the independent variable, this probability is equal to 0.5 (i.e. the distribution of the regression coefficients is centred on zero). All comparative analyses were performed using the software R, through packages ‘APE’ [19] and ‘bmaMCMCAnalysis’ (L. Spangenberg, R. Romero and H. Naya; available upon request). Phylogenetically informed analyses were conducted only for the selected ‘good models’ for practical reasons, i.e. the inability to run the phylogenetic analyses for all the 57,343 models, given the computational costs. Thus, it is possible that phylogenetic ‘good models’ are not among the set of non-phylogenetic ‘good models’ that were considered in the phylogenetic analyses.

(c) Data analysis: relationship between \(C_{\text{min}}\) and basal metabolic rate

To evaluate whether there is compensation between \(C_{\text{min}}\) and BMR, we first estimated the value of the slope between \(\log_{10}(C_{\text{min}})\) and \(\log_{10}(\text{BMR})\) through a standard least-squares regression analysis, using \(\log_{10}(m_b)\) as a covariate. Then, we tested whether the value of this slope was statistically different from one, using a two-tailed Student’s t-test. In addition, an analysis of statistical power was performed in order to determine the level of confidence around the null hypothesis (i.e. slope = 1). On the other hand, given that for the analysed dataset, \(T_{\text{med}}\) is the best single predictor of mass-independent BMR [20] and also mass-independent \(C_{\text{min}}\) (see results), we used a structural equation modelling approach [21] to compare three causal models: (i) annual mean temperature independently affects BMR and \(C_{\text{min}}\) \((T_{\text{med}} \rightarrow \log_{10}\text{BMR}, T_{\text{med}} \rightarrow \log_{10}C_{\text{min}})\), (ii) mean annual temperature directly affects BMR and indirectly \(C_{\text{min}}\) \((T_{\text{med}} \rightarrow \log_{10}\text{BMR} \rightarrow \log_{10}C_{\text{min}})\) and (iii) mean annual temperature directly affects \(C_{\text{min}}\) and indirectly BMR \((T_{\text{med}} \rightarrow \log_{10}C_{\text{min}} \rightarrow \log_{10}\text{BMR})\). In all these models, body mass \((\log_{10}m_b)\) was included as an exogenous variable affecting both \(C_{\text{min}}\) and BMR. We used a maximum-likelihood method to estimate the general fit of each model as well as model parameters. The significance of each model was assessed with \(\chi^2\) statistics, which compare the fit between the observed and predicted elements of the covariance matrix [21]. A significant \(\chi^2\)-value means that the tested model is not supported by the data. All these analyses were performed using the modules.
3. Results

(a) Exogenous factors affecting minimal thermal conductance

Statistical analyses with only one independent variable indicate that annual mean temperature was (by far) the best single predictor of mass-independent \( C_{\text{min}} \) (figure 1a and table 1). In line with this, multiple factor analyses indicate that models including annual mean temperature plus annual accumulated rainfall/aridity were the best models to explain variation in mass-independent \( C_{\text{min}} \) (table 2; electronic supplementary material, table S2). A third multiple factor model including latitude, in addition to annual mean temperature and annual accumulated rainfall, could also be included in the group of ‘good models’, but its \( \Delta \text{BIC} \) value was close to the limit of 2.3 units (table 2; electronic supplementary material, table S2). Regarding phylogenetic analyses, all independent variables—except latitude—had a significant \((p < 0.05)\) or marginally significant \((p < 0.1)\) effect on mass-independent \( C_{\text{min}} \) (table 3).

(b) The relationship between \( C_{\text{min}} \) and basal metabolic rate

After controlling for the effect of body mass, there was a strong positive correlation between \( \log_{10}(C_{\text{min}}) \) and \( \log_{10}(\text{BMR}) \) (figure 1b). In addition, the slope of this correlation \((X \pm 1 \text{ s.e.} = 1.0987 \pm 0.0651)\) was not statistically different from one \((t_{124} = 1.51, p > 0.13)\), indicating an almost perfect compensation between the two energetic variables. However, it is important to note that, with our dataset, the probability of rejecting the null hypothesis (i.e. slope \( = 1 \)) is larger than 0.80 just for deviation larger than 0.18 from a slope value of one. Finally, results from structural equation modelling indicated that the model where \( C_{\text{min}} \) values were adjusted to BMR values (model no. 2) was the only model supported by the data (figure 2).

4. Discussion

Global assessment of physiological variables is of paramount relevance in the current scenario of climate change. After all, if we intend to predict how species will respond to accelerated human-caused changes in environmental factors, we need to understand how they respond to these factors in an ecological and evolutionary time frame. Several studies analysing global scale variation in physiological variables have been published in recent years, mainly focusing on metabolic rates [10,13,22–26], metabolic scopes [27,28], thermal tolerance ranges [29–32] and physiological flexibility [33–35]. However, only minor efforts have been made to analyse other relevant physiological variables; such is the case of thermal conductance (but see [10]).

Three results obtained in this study are noteworthy. First, we quantitatively demonstrate that some climatic variables are correlated with thermal conductance. Specifically, single factor statistical models indicate that annual temperature is the best single predictor of mass-independent \( C_{\text{min}} \), while...
multiple factor statistical models indicate that annual mean temperature plus annual accumulated rainfall/aridity are the major determinants of mass-independent $C_{\text{min}}$. Interestingly, a recent analysis of 195 rodent species indicates that environmental temperatures ($T_{\text{med}}$ or $T_{\text{min}}$) plus accumulated rainfall (or NPP) are the main predictors of mass-independent variation in BMR [20]. Thus, it appears that the same exogenous factors are affecting the evolution of both $C_{\text{min}}$ and BMR. Strikingly, our results do not agree with those reported by Lovegrove [10], who found that the coefficient of variation of annual mean rainfall was the only climatic variable correlated with $C_{\text{min}}$. A potential explanation for this discrepancy is that Lovegrove [10] analysed mass-specific $C_{\text{min}}$ values (i.e. $C_{\text{min}}/m_b$) instead of overall $C_{\text{min}}$ values. To evaluate this possibility, we re-run

| Parameter estimation ($B$), standard deviation (s.d.) and proportion of posterior estimates greater than zero (gt0) for each independent variable included in each selected model (table 2), according to phylogenetically informed analysis. See material and methods for exogenous factors abbreviations. Body mass contribution was highly significant in all cases (gt0 > 0.99). |

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Figure 2. Graphical representation of the three causal models evaluated, together with the associated statistics and the partial correlation coefficient for each pathway ($\pm$ 1 s.e.). $T_{\text{med}} = $ annual mean temperature. The log10 of BMR, $C_{\text{min}}$ and $m_b$ were used in this analysis. Total number of observations = 127. p-values were smaller than 0.01 for all partial correlations (except for those between $T_{\text{med}}$ and $m_b$).

Figure 3. (a–c) Three hypothetical scenarios for seasonal changes in BMR and $C_{\text{min}}$ (i.e. the slope of the metabolic curves below the thermoneutral zone) in a tropical and a temperate rodent species (see text for a detailed explanation). Note that current data appear to support the scenario depicted in (a). (Online version in colour.)
single factor analyses using $C_{\min}/m_b$ values (data not shown) and found that aridity was the only factor correlated with mass-specific thermal conductance ($p < 0.05$), which is fairly congruent with Lovegrove’s former conclusion. Second, we note that—as could be expected from the Scholander–Irving classical model of heat transfer [3,4]—there is an almost perfect compensation between BMRs and $C_{\min}$. This result reinforces the idea that the two physiological variables are part of a coordinated system for heat transfer that affects body temperature regulation in endothermic animals [5]. Third, structural equation modelling indicates that annual mean temperature directly affects BMR, and that (summer) $C_{\min}$ values are then adjusted (in an evolutionary sense) to changes in BMR values.

Interestingly, our results are congruent with recent ideas about the evolution of BMR, such as the ‘obligatory heat’ model (OHM) [20,27] and the ‘heat dissipation limit’ theory (HDLT) [36,37]. For instance, it could be possible that colonization of colder environments (e.g. after glacial periods) favours an increase in BMR through a rise in the size of metabolically expensive organs (as stated by the OHM); but, at the same time, this determined an increase in thermal conductance during the warmer months of the year (as shown here) in order to avoid overheating (as suggested by the HDLT) (figure 3a). In line with this, it is noteworthy that: (i) circumstantial evidence suggests that small tropical and arctic mammals collected during winter appear to be equally well insulated [2,3]; (ii) in contrast with mean and minimum temperatures, maximum temperatures in terrestrial environments change slightly with latitude across the globe [10,29]. In any case, if the scenario described above (and depicted in figure 3a) is correct, a remaining question is why species inhabiting high-latitude environments ‘choose’ to have larger BMR values all year round and increase thermal conductance during summer, instead of (i) maintaining constant BMR and reducing thermal conductance during winter (figure 3b) or (ii) maintaining constant thermal conductance and reducing BMR during summer (figure 3c). Regarding the first alternative, a potential answer is that reducing thermal conductance during winter may represent a costly process that does not result in other benefits than thermal regulation, while maintaining high BMR values—through adjustments in organs size—may represent a costly process that results in several other benefits (e.g. an ability to exploit low-quality diets during winter) in addition to thermal regulation [20]. Another potential explanation for this alternative is that minimum conductance values occurring at low latitudes (i.e. before the colonization of cold habitats) are close to the minimum possible—owing to, for example, an impaired ability to move with a very large fur coat—and thus, a further reduction in thermal conductance is not feasible. With regard to the second alternative, it is fairly clear that a reduction in BMR during summer months—i.e. when food is abundant—may not represent a good strategy, because BMR is positively correlated with reproductive output [5,38,39].

Finally, two additional points should be noted. First, our analysis is based on interspecific comparisons at a global scale; it does not deny that intraspecific adjustments in BMR can occur [40] but just implies that intraspecific changes are of lower magnitude than interspecific variation in mean values (as it appears to be the case, e.g. compare [20] with [41]). Second, if all the ideas mentioned above are correct, it should be expected that seasonal change in thermal conductance (i.e. a measure of phenotypic flexibility) should increase with latitude, as predicted by the climatic variability hypothesis [40,42].

To end, we care to mention a caveat about the generality of our findings. $C_{\min}$ in small-sized mammals corresponds with the lower limit of thermoneutrality, and hence, $C_{\min}$ values have a clear interpretation from the Scholander–Irving heat transfer model. This may not be the case for intermediate and large animals, in which a circulatory separation between core and shell temperature could determine a reduction in thermal conductance values beyond those observed at the lower limit of thermoneutrality [5,43]. Undoubtedly, more studies analysing physiological diversity in different taxonomic groups and taking into account large geographical and temporal scales are needed to understand and predict human impacts on the Earth’s ecosystems [44].

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